

# Using genome-wide complex trait analysis to quantify ‘missing heritability’ in Parkinson’s disease

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**Genome-wide association studies (GWASs) have been successful at identifying single-nucleotide polymorphisms (SNPs) highly associated with common traits; however, a great deal of the heritable variation associated with common traits remains unaccounted for within the genome. Genome-wide complex trait analysis (GCTA) is a statistical method that applies a linear mixed model to estimate phenotypic variance of complex traits explained by genome-wide SNPs, including those not associated with the trait in a GWAS. We applied GCTA to 8 cohorts containing 7096 case and 19 455 control individuals of European ancestry in order to examine the missing heritability present in Parkinson’s disease (PD). We meta-analyzed our initial**

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**results to produce robust heritability estimates for PD types across cohorts. Our results identify 27% (95% CI 17–38,  $P = 8.08E - 08$ ) phenotypic variance associated with all types of PD, 15% (95% CI –0.2 to 33,  $P = 0.09$ ) phenotypic variance associated with early-onset PD and 31% (95% CI 17–44,  $P = 1.34E - 05$ ) phenotypic variance associated with late-onset PD. This is a substantial increase from the genetic variance identified by top GWAS hits alone (between 3 and 5%) and indicates there are substantially more risk loci to be identified. Our results suggest that although GWASs are a useful tool in identifying the most common variants associated with complex disease, a great deal of common variants of small effect remain to be discovered.**

## INTRODUCTION

Genome-wide association studies (GWASs) have been successful at identifying single-nucleotide polymorphisms (SNPs) highly associated with common traits; however, a great deal of the heritable variation associated with common traits remains unknown (1–3). Parkinson's disease (PD) is the second most common neurodegenerative disease, and is clinically characterized by rigidity, resting tremor and bradykinesia (slowed movement). The average age at onset is 68 years, however this is highly variable and can range from adolescence to old age (4–6). Generally, individuals with disease onset before age 55 are categorized as early onset, and those with disease onset after age 55 are categorized as late onset (7–9).

GWASs have identified risk variants at over two dozen loci influencing PD risk; however, these are thought to explain only a fraction of the variance in PD liability (8,10). In addition, mutations known to cause monogenic forms of PD account for only a small proportion of disease (11). These include mutations in  $\alpha$ -synuclein (*SNCA*), parkin (*PARK2*), DJ-1 (*PARK7*), PTEN-induced putative kinase I (*PINK1*) and leucine-rich repeat kinase 2 (*LRRK2*) (12–14). Although additional loci that have been implicated in PD risk produce small effects, collectively they comprise a larger portion of genetic component responsible for the development of PD (15).

Twin studies are useful in differentiating the impacts of genetics and the environment as the sources of a disease. If genetic factors greatly influence the presence of a disease, it is expected that concordance in monozygotic (MZ) twins will be greater than dizygotic (DZ) twins. Twin studies of PD have shown very low pairwise concordance, estimated as  $\sim 0.129$ , with no discernible difference between MZ and DZ pairs, although there is some concern over the cross-sectional nature of these studies. In addition, concordance-adjusted prevalence in twin studies has been estimated at 8.67/1000, or  $< 0.1\%$  (16). This is a similar prevalence value to what is seen in other population-based studies of PD prevalence (4,8,17,18).

Monogenic familial forms of PD are often early-onset, and it has been suggested that early-onset PD has a greater familial/heritable component in its etiology than later onset manifestations of the disease (19). It is likely that the etiology of early-onset PD is dissimilar to late-onset PD, which seems to occur more sporadically (15). At least one twin study has performed analyses controlled for age, with results showing significant rates of concordance for MZ pairs with early age at onset of PD, and a near lack of concordance for later onset individuals. This strongly supports the role of genetics

in early-onset PD; however, the sample size for this subset of individuals was very small, and shared environmental exposure in twins cannot be overlooked (16). Furthermore, a longitudinal twin study examined both MZ and DZ pairs over an 8-year period and observed a significantly higher concordance rate among MZ pairs; over time, concordance in these individuals increased from 0 to 33% (20). These data suggest an additional environmental trigger in the etiology of PD and further demonstrate that a simple genetic model does not fit the etiology of this disease (16,20,21).

We sought a method to capture the genetic variance associated with PD. Genome-wide complex trait analysis (GCTA) estimates the components of phenotypic variance, that is, the polygenic additive variance (heritability) that can be explained by genome-wide SNPs, including those that are not significantly associated with the phenotype of interest in GWASs (22). Notably, this method provides a lower limit for trait heritability because it is improbable that all causal variants have been exactly tagged by the SNPs on the genotyping platforms used to perform initial analyses (15). GCTA works by utilizing genome-wide SNPs to quantify the phenotypic variance of all putative causal variants, as opposed to only genome-wide significant trait-associated variants. For a given disease trait, the heritability explained by all SNPs is estimated simultaneously, as opposed to testing the association of some previously identified SNPs from GWASs, which relies on shared linkage disequilibrium across common variants associated with disease status to quantify heritability (1–3).

Using GCTA, we aim to identify a larger segment of additive genetic variance not typically associated with early- and late-onset PD via GWASs in central European and Scandinavian ancestry populations, thus generating the most comprehensive estimates to date of PD heritability (23). Recent analyses have also used GCTA to examine the dichotomy of early and late-onset PD, using self-reported diagnosis data collected by the personal genetics company 23andMe (15). Do *et al.* (15) report estimates for early-onset PD heritability as  $\sim 0.306$  (95% CI 0.136–0.476) and for late onset as  $\sim 0.285$  (95% CI 0.2224–0.346). These estimates are supported by our meta-analysis, which provides additional support for the complex genetic structure of PD.

Our analyses offer a continuation and extension of this previous research by dichotomizing the PD phenotype based on age and subsets of SNPs; in addition, we imputed our data to test both typed and untyped variants when estimating phenotypic variance of the PD phenotype. This not only encapsulates more putatively associated variants than previous studies yielding a more informed estimate, but it is also in line

with the methodologies that will be used in future GWAS studies. Critically, imputation is useful in this study because of the inclusion of multiple clinically derived studies as opposed to a single cohort study of self-reported cases. In particular, the use of imputation provided a means of facilitating more cohesive data for meta-analyses, as our genotyped data were standardized and expanded by the imputation of millions of shared SNPs per cohort. By imputing genotypes across cohorts, we are able to test the same set of SNPs regardless of the limitations of the initial platform.

## RESULTS

Cohort-level heritability estimates are provided in Table 1 and accompanying forest plots are shown in Figure 1. Our heritability estimates vary from ~16% (UK, SE = 0.027) to ~49% (Finland, SE = 0.04). Moderate estimates were obtained for the majority of cohorts.

Results of our meta-analysis are presented in Table 2. We generated statistically significant heritability estimates for all PD types and late-onset PD across all SNP sets at the meta-analysis level. Our results identify 27% (95% CI 17–38,  $P = 8.08E - 08$ ) phenotypic variance associated with all PD samples, 15% (95% CI -0.2 to 33,  $P = 0.09$ , non significant) phenotypic variance associated with early-onset PD and 31% (95% CI 17–44,  $P = 1.34E - 05$ ) phenotypic variance associated with late-onset PD. These estimates are a substantial increase from the genetic variance identified by GWA top SNPs alone (3–5%). The estimates for all PD and late-onset PD types increased moderately using imputed data (estimate using only genotyped SNPs—all samples: 24%,  $P = 1.33E - 07$ ; estimate using only genotyped SNPs—late-onset samples: 26%,  $P = 7.34E - 07$ ); however, these estimates are not significantly different and are attributed to unexplained variation. The early-onset estimate decreased drastically and lost significance when imputed data were used (estimate using only genotyped SNPs—early-onset samples: 33%,  $P = 3.91E - 04$ ; see Table 3 for additional detail). Despite the apparent difference between genotyped and imputed early-onset samples, the reported estimates remain within the 95% confidence interval ranges of each data type.

The data suggest a large portion of heritability in PD has not yet been accounted for in current GWASs. Although a small degree of heritable genetic variation is tagged by common SNPs (i.e. those included on micro-array genotyping assays), recent studies point toward the role of rare variants in disease etiology (24–26). It was expected that combined heritability estimates would show early-onset PD as having a higher portion of phenotypic variance than late-onset PD; however, this was not observed in the pooled analyses or the meta-analysis. The remaining variability in PD etiology not ascribed to heritable factors in our analyses suggests a contribution of combined effects of rare variants not tagged by current genotyping, possible environmental factors or a stochastic component. In addition, other factors including non-additive genetic effects and artifacts in the SNP data may also contribute to this discrepancy. We noted a priori 3 263

728 of rare variants were of lower quality and poorly tagged by the available microarray genotype data (57.69%).

## DISCUSSION

Our estimates of phenotypic variance provide unequivocal and compelling evidence of yet-to-be-discovered additional genetic factors that contribute to the etiology of PD. While imputation can capture up to ~50% of the genetic variation associated with PD in some cohorts, our study was limited by the sensitivity of microarray-based genotyping methods utilized in our calculation of heritability. In addition, disparate demographic histories between cohorts likely contributed to the differences shown in the cohort-level heritability estimates. In particular, the Icelandic cohort produced the highest genome-wide estimates and the lowest GWAS SNPs estimates. The population of Iceland is remarkably homogenous compared with the populations of France and the UK, for example, and their shared ancestry is reflected in their genetic structure as measured by GCTA.

Our results also provide support for the hypothesis that rare variants of potentially large effect are less likely to be accurately tagged by microarray, potentially biasing heritability estimates as appearing lower than what should be expected, particularly for early-onset PD. Large-scale genome and exome sequencing in conjunction with denser genotyping in large cohorts may help to better quantify the heritability of complex diseases. These efforts will also aid in the identification of loci that contribute to ‘missing heritability’ previously undetected in earlier generation technologies used for capturing genomic variation, as were implemented in this study. For example, the GBA locus contains approximately 17 rare mutations, not all of which were originally detected by GWASs (27,28). In addition, the use of newer technologies is likely to revise disease heritability estimates upward.

Heterogeneity in the heritability estimates reported here could result from heterogeneity in the coverage of the genome as well as from patient acquisition biases. A large proportion of the dbGaP cohort (herein referred to as MF, further described in Materials and Methods) is comprised of familial PD cases; however, the heritability estimates for this cohort are among the lowest. This may reflect differences in genotyping platform. The US-NIA (US), Dutch (NL), UK: WTCCC2/Cardiff (UK), French (FR) and German (GER) cohorts were genotyped using the 610 and 550K Illumina arrays containing 500K SNPs, whereas the MF, Finnish (FIN) and Icelandic (ICE) cohorts were genotyped using the Illumina 370K. The use of MACH (Markov Chain-based haplotyper) for imputation allows comparisons between cohorts of different genotyping platforms to be made, because genotypes are imputed based on the observed haplotype structure of each cohort, and therefore analysis occurs across the same set of SNPs.

It is important to note that the choice of prevalence value has a minor impact on our heritability estimates. Our study model dichotomizes the PD phenotype based on age, and because PD prevalence increases with age, we used a prevalence value standardized for age and gender, specified here as 0.002. To account for the larger prevalence within the

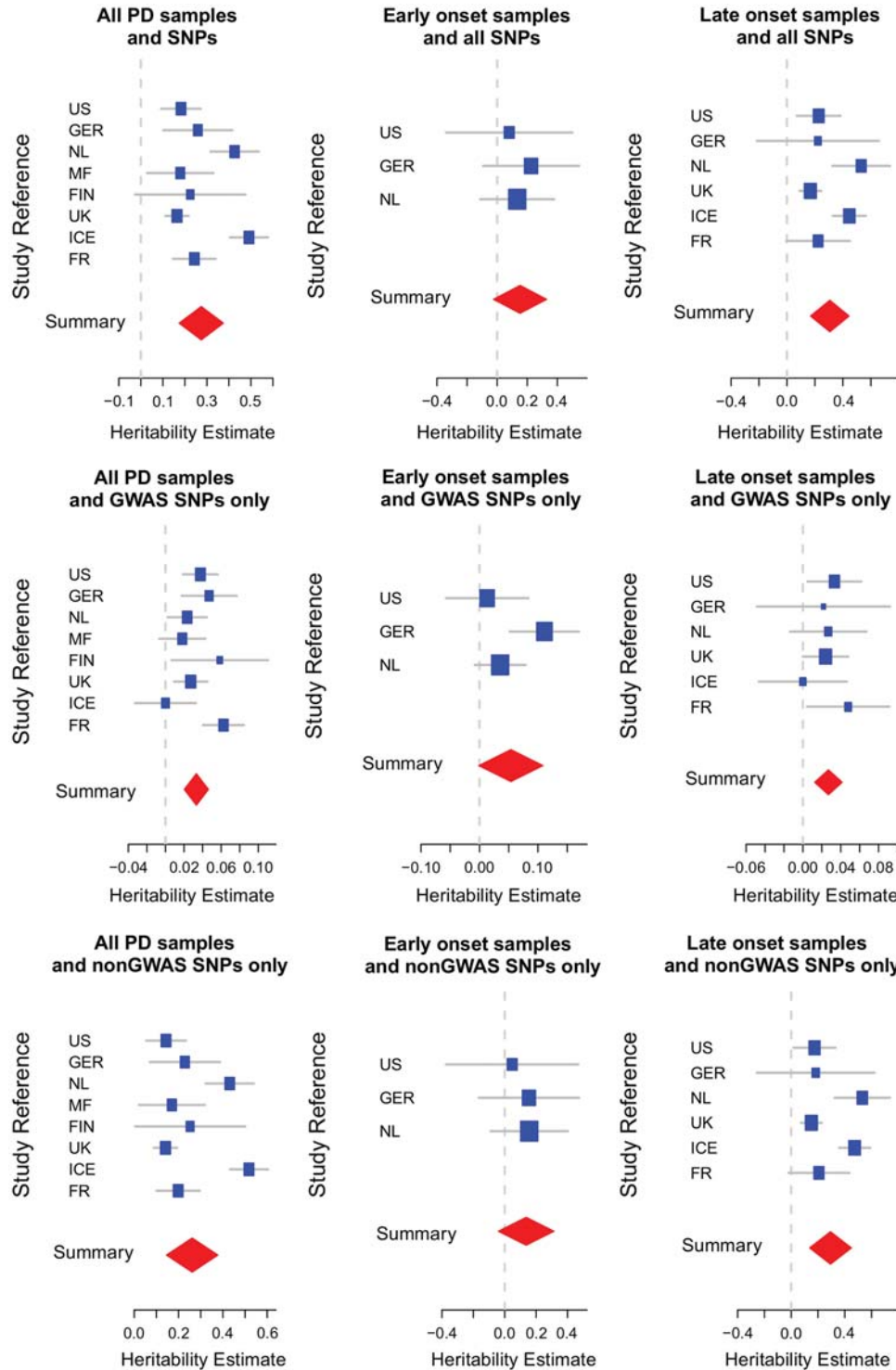
**Table 1.** Cohort level analyses

		Heritability estimate	SE of heritability estimate	Cases ( <i>n</i> )	Controls ( <i>n</i> )
Cohort: US					
All samples	All SNPs	0.182545	0.045998	971	3034
	PD GWAS SNPs in PD loci/regions	0.037605	0.009569	971	3034
	PD GWAS regions excluded	0.143514	0.045902	971	3034
Early onset	All SNPs	0.080189	0.214376	365	276
	PD GWAS SNPs in PD loci/regions	0.013059	0.035854	365	276
	PD GWAS regions excluded	0.047529	0.214932	365	276
Late onset	All SNPs	0.227307	0.080002	572	1620
	PD GWAS SNPs in PD loci/regions	0.033377	0.014531	572	1620
	PD GWAS regions excluded	0.174017	0.079982	572	1620
Cohort: GER					
All samples	All SNPs	0.259042	0.080298	742	944
	PD GWAS SNPs in PD loci/regions	0.047032	0.015157	742	944
	PD GWAS regions excluded	0.228625	0.0807	742	944
Early onset	All SNPs	0.22548	0.162678	302	670
	PD GWAS SNPs in PD loci/regions	0.111	0.030216	302	670
	PD GWAS regions excluded	0.154317	0.162772	302	670
Late onset	All SNPs	0.221245	0.222233	367	267
	PD GWAS SNPs in PD loci/regions	0.021783	0.035878	367	267
	PD GWAS regions excluded	0.183247	0.22393	367	267
Cohort: NL					
All samples	All SNPs	0.426036	0.055851	772	2024
	GWAS SNPs in PD loci/regions	0.02337	0.010787	772	2024
	PD GWAS regions excluded	0.430828	0.055468	772	2024
Early onset	All SNPs	0.133057	0.125479	366	871
	GWAS SNPs in PD loci/regions	0.035313	0.022193	366	871
	PD GWAS regions excluded	0.156223	0.125421	366	871
Late onset	All SNPs	0.531087	0.105094	379	1148
	GWAS SNPs in PD loci/regions	0.026801	0.02065	379	1148
	PD GWAS regions excluded	0.531087	0.105496	379	1148
Cohort: UK					
All samples	All SNPs	0.164206	0.027159	1705	5200
	GWAS SNPs in PD loci/regions	0.027181	0.009195	1705	5200
	PD GWAS regions excluded	0.141085	0.027061	1705	5200
Late onset	All SNPs	0.167639	0.039548	1258	2699
	GWAS SNPs in PD loci/regions	0.023974	0.012248	1258	2699
	PD GWAS regions excluded	0.150867	0.039547	1258	2699
Cohort: ICE					
All samples	All SNPs	0.491238	0.044370	604	4916
	GWAS SNPs in PD loci/regions	0.000001	0.016722	604	4916
	PD GWAS regions excluded	0.517731	0.044039	604	4916
Late onset	All SNPs	0.446607	0.060354	449	2624
	GWAS SNPs in PD loci/regions	0.000001	0.023778	449	2624
	PD GWAS regions excluded	0.473371	0.059609	449	2624
Cohort: FR					
All samples	All SNPs	0.243203	0.04944	1039	1984
	GWAS SNPs in PD loci/regions	0.062499	0.011242	1039	1984
	PD GWAS regions excluded	0.198351	0.049936	1039	1984
Late onset	All SNPs	0.222832	0.116052	340	1984
	GWAS SNPs in PD loci/regions	0.047813	0.022223	340	1984
	PD GWAS regions excluded	0.206995	0.116277	340	1984
Cohort: MF					
All samples	All SNPs	0.179096	0.077089	876	857
	GWAS SNPs in PD loci/regions	0.018118	0.01266	876	857
	PD GWAS regions excluded	0.170097	0.076168	876	857
Cohort: FIN					
All Samples	All SNPs	0.224504	0.127949	387	496
	GWAS SNPs in PD loci/regions	0.058549	0.026548	387	496
	PD GWAS regions excluded	0.252043	0.126645	387	496

Cohort-level heritability estimates from imputation.

case population and to control for ascertainment, we used GCTA to transform the explained variance estimate from the observed scale ( $V(1)/V_p$ ) to the underlying scale ( $V(1)/V_{p\_L}$ ). Previous work has shown that assumed prevalence has only a small impact on GCTA estimates. In particular,

Do *et al.* (15) show that prevalence values ranging 3-fold (0.005–0.015) impact on their estimates by only ~5%. Our subsequent analyses show that using a conservative prevalence value still provides ample evidence for identifying increased phenotypic variance of PD.



**Figure 1.** Forest plots of heritability estimates across cohorts. Cohort-specific heritability estimates are shown in blue, the size of the square is proportional to the size of the study. Confidence intervals of the summary heritability estimates are shown as red diamonds, with the centerline of each diamond representing the summary heritability estimate for that particular subset of data.

An alternative method of analysis would employ a conditional analysis within GCTA, using the GWAS-significant PD SNPs as covariates to statistically correct for known signals during the restricted maximum likelihood (REML)

analysis. It is expected that the known SNPs would capture the same variance as the whole set of SNPs used in the GWAS region analysis; however, three of the top SNPs (rs1491942, rs6710823 and rs76763715) were not available

**Table 2.** Summary of random-effects meta-analysis from imputation

PD type	SNPs included in analysis	Heritability estimate from random effects	Lower 95% confidence interval	Upper 95% confidence interval	<i>P</i> -value from random effects	Heterogeneity of variance from random effects (%)	Heterogeneity <i>P</i> -value
All	All SNPs	0.27	0.17	0.38	8.80E - 08	0.02	0.00E + 00
	GWAS SNPs in PD loci/regions	0.03	0.02	0.05	5.23E - 07	0.00	3.20E - 02
	Non-GWAS regions	0.26	0.15	0.38	9.69E - 06	0.02	0.00E + 00
Early onset	All SNPs	0.15	-0.02	0.33	9.16E - 02	0.00	8.44E - 01
	GWAS SNPs in PD loci/regions	0.05	0.00	0.11	5.50E - 02	0.00	6.20E - 02
Late onset	Non-GWAS regions	0.14	-0.04	0.31	1.30E - 01	0.00	9.01E - 01
	All SNPs	0.31	0.17	0.44	1.34E - 05	0.02	1.00E - 03
	GWAS SNPs in PD loci/regions	0.03	0.01	0.04	2.59E - 04	0.00	7.85E - 01
	Non-GWAS regions	0.29	0.14	0.45	2.30E - 04	0.03	0.00E + 00

Meta-analysis of heritability estimates from imputed data. Results are significant for All and late-onset subset of PD.

**Table 3.** Summary of random-effects meta-analysis from genotyping

PD type	SNPs included in analysis	Heritability estimate from random effects	Lower 95% confidence interval	Upper 95% confidence interval	<i>P</i> -value from random effects	Heterogeneity of variance from random effects (%)	Heterogeneity <i>P</i> -value
All	All SNPs	0.24	0.16	0.31	6.27E + 00	0.01	4.00E - 03
	GWAS SNPs in PD loci/regions	0.02	0.01	0.03	1.77E - 04	0.00	3.30E - 02
	Non-GWAS regions	0.23	0.14	0.32	5.05E - 07	1.30	<1.00E - 16
Early onset	All SNPs	0.33	0.15	0.52	3.91E - 04	0.00	7.45E - 01
	GWAS SNPs in PD loci/regions	0.02	-0.01	0.05	1.38E - 01	0.00	5.94E - 01
Late onset	Non-GWAS regions	0.32	0.13	0.50	7.33E - 04	0.00	6.45E - 01
	All SNPs	0.26	0.16	0.36	7.34E - 07	1.00	2.10E - 02
	GWAS SNPs in PD loci/regions	0.01	0.00	0.02	1.98E - 02	0.00	5.63E - 01
	Non-GWAS regions	0.25	0.15	0.35	9.11E - 07	1.00	2.70E - 02

Meta-analysis of heritability estimates from genotyped data.

**Table 4.** Study descriptives

Cases Cohort	Cases			Controls			Study details	
	Sample size	Female (%)	Mean age at onset [years (SD)]	Sample size	Female (%)	Mean age [years (SD)]	Number of SNPs used for imputation	Genomic inflation factor ( $\lambda$ )
US	937	40.23	57.81 (13.16)	3033	52.82	63.3 (10.06)	545 066	1.035
GER	740	39.59	49.33 (22.21)	944	47.99	47 (13.25)	561 467	1.025
NL	771	36.45	53.39 (16.62)	2024	55.94	55.56 (6.60)	546 155	1.061
UK	1705	42.22	48.95 (13.84)	5200	49.42	53 (0)	532 616	1.034
ICE	604	47.85	73.26 (13.84)	5520	55.87	85.12 (10.77)	316 905	1.011
FR	1039	41.2	48.9 (12.8)	1984	42.9	73.7 (5.4)	492 929	1.03
MF	876	40.41	36.42 (11.08)	857	60.21	NA	325 770	1.013
FIN	387	45.99	48.28 (6.97)	496	0	91.98 (7.46)	302 463	1.066

Cohort-level descriptive statistics and study details.

in our imputed data for multiple cohorts. As no proxies were available, the results are not comparable with the unconditional analysis and are not reported here.

Our analyses indicate that the estimates of heritability presented here are minimal estimates, despite the significant increase over what has been identified by GWASs. The genetic additive heritability identified by the GWAS SNPs is also likely an underestimate of the genetic variance due to common PD variants, as the associated GWAS SNPs are not necessarily causal: their effects are lower than those of the true susceptibility variants. Genotyping these variants would likely produce greater estimates of the common genetic additive variance due to the GWAS PD loci. Indel and structural variants were not considered here, although a portion of these was likely tagged by the considered SNPs. In addition, the portion of variance explained by GWAS SNPs is underestimated by GCTA, as the model places a prior centered zero as the effect size of the SNPs that are used in the calculation of the genetic relationship matrix (GRM). This has been previously shown using sparse regression techniques, which account for 6–7% of total variance in GWAS SNPs, further suggesting the majority of genetic variants that contribute to PD lay outside of GWAS-significant SNPs (15). Large-scale sequencing efforts often identify many variants that are not easily tagged using microarray genotyping. This is apparent in evaluations of the success of capturing rare variants, most commonly defined as having a minor allele frequency (MAF) of  $\leq 0.05$ . In addition, rare variants are thought to harbor larger deleterious effects than more common variants (29,30). In this context, it is likely that further GWA studies that include greater numbers of typed SNPs will identify additional risk for PD.

In conclusion, our estimates do not confirm hypotheses of a greater contribution of genetic risk for early-onset PD compared with late-onset PD; however, we expect the contributions of rare variants not tagged by microarray genotyping to have a substantial impact on the genetic contributions to disease risk. Although our imputed analysis did not provide significant results for this subset of PD, we note a priori that a significant portion of rare variants are not well captured in microarrays (~58%); therefore, our heritability estimates are lower than what might be expected. In addition, very rare variants (MAF < 0.005) require reference panels with at least 1200 subjects in order to impute, as it is necessary to

observe multiple copies of an SNP in order to accurately constitute the haplotypes that will be used for imputation (31). Future analyses incorporating denser genome and exome-wide assays in conjunction with newer sequencing technologies will likely see increased heritability estimates associated with PD and other complex traits, as a significantly larger genetic contribution to disease risk is identified.

## MATERIALS AND METHODS

### Study populations

We utilized six GWAS data sets drawn from the International Parkinson's Disease Genomics Consortium (IPDGC), comprised of US and European participants previously described in detail elsewhere (7,32–37). The IPDGC includes data from Iceland (ICE,  $n = 5520$ ), the UK: WTCCC2/Cardiff (UK,  $n = 6905$ ), the Netherlands (NL,  $n = 2796$ ), Germany (GER,  $n = 1686$ ), France (FR,  $n = 3023$ ) and the USA-NIA (US,  $n = 4005$ ). Additional US participant data were obtained via dbGaP, made available by the NINDS Human Genetics Resource Center and have been described by Pankratz *et al.* (35) (MF,  $n = 1733$ ); a Finnish case–control cohort was also incorporated into this study based on the Vantaa 85+ Study (38) (FIN = 883; cases from Oulu,  $n = 387$ ; controls from mitoPARK and Vantaa 85+,  $n = 496$ ). In total, 8 cohorts comprising 7096 cases and 19 455 controls were used. Study descriptives for each cohort are outlined in Table 4. Cohorts containing individuals with age data were further subset to account for early- and late-onset cases and controls. Age at onset was defined here as age at initial diagnosis in cases and matched to controls using age at study ascertainment. Early onset was quantified as  $\leq 55$  years old and late onset was quantified as  $> 55$  years old. Early-onset data sets were drawn only from individuals in the US, German and Dutch cohorts, as the remainder of the cohorts did not contain control individuals  $\leq 55$  years of age. Late-onset data sets were drawn from individuals in the US, German, Dutch, Icelandic, French and UK cohorts.

As per our study design, older controls were excluded from our early-onset analyses due to previous research suggesting demographic factors such as age influence genetic substructure. In particular, increasing homozygosity has been associated with chronological age (39). As this trend could

influence the frequency of the variants used in our analyses of early-onset PD heritability and introduce stochastic change or other potential bias that is not controlled for, we have excluded older controls from early-onset analyses.

Stringent quality control measures were applied to all data sets prior to GCTA analysis, in order to control for ascertainment bias and any artifacts introduced into the data by the genotyping process (32). It was necessary to employ more stringent quality control than what is generally acceptable for GWASs because systematic differences between cases and controls could be picked up as genetic variance (3). Cases and controls in all cohorts were analyzed together to reduce bias. SNPs with different call rates in cases and controls were excluded. Additional detail regarding specific cohort-level quality control measures is described elsewhere (32,40).

As a note, three SNPs, *ACMSD*, *GAK*, and *HLA-DRB5*, were initially found to be associated in a GWAS that was performed on the same data sets used in the present study (32). Known inherited Mendelian loci were not directly accounted for in our estimates.

### Quality control and imputation

After samples were collected from each cohort, and standard quality control was applied, MACH (version 1.0.16) was applied to each cohort to impute genotypes for all European ancestry participants. Haplotypes were derived from 500 European ancestry samples in the 1000 Genomes Project (as of 23 September 2011), based on initial low coverage and exome sequencing. A quality threshold of a minimum 0.30 squared correlation between proximal, experimental and imputed genotypes was applied, indicated by the  $R^2$  metric from MACH (41,42). We used the default settings of MiniMac (41,42) to impute variants into each cohort, and the total number of imputed variants per cohort is shown in Table 4.

Data were imputed in a two-stage design. The first stage of imputation generated error and crossover maps on a random subset of 200 samples per study, for 100 iterations of the statistical model. These maps were used as parameter estimates for imputation, to generate maximum likelihood estimates of allele dosages per SNP, on the basis of reference haplotypes from the 1000 Genomes Project during the second stage of the imputation. SNPs were excluded if their  $R^2$  quality estimates were  $<0.30$ , as estimated by MACH, or if their MAFs were  $<0.01$ , because imputed genotypes below these values are likely poor in quality and are more susceptible to errors in imputation.

Cryptic relatedness among samples was addressed at both cohort and meta-analysis levels. Within the cohorts, samples sharing  $>0.15$  proportion of alleles or samples identified as first or second degree relatives according to identity by descent estimates were excluded (32). As an additional quality control measure, we calculated genomic control for individual data sets. Genomic control is based on the  $\chi^2$  statistic and is estimated as the deviance of the median test statistic distribution from the expected null. The product of this analysis is a  $\lambda$ -value;  $\lambda$ -values  $<1.05$  are standard in GWASs (32).  $\lambda$ -Values were obtained before imputation, and indicate that

population stratification is minimal. Genome-wide  $\lambda$ -values for each cohort are reported in Table 4.

### Statistical analysis

We applied GCTA to estimate heritability within each of the stratified data sets per cohort. Heritability is defined here as the proportion of phenotypic variation in a population that is due to genetic variance between individuals. GRMs were calculated for each subset of data to determine the genetic relationship between pairs of individuals (3,22). The GRMs in this analysis were estimated using imputed dosage score SNP data; therefore, the estimation of variance explained by the SNPs relies on the  $R^2$  cutoff used to select the SNPs.

GRMs were input into an REML analysis to produce estimates of the proportion of phenotypic variance explained by the SNPs within each subset of data (22). In addition to cohort-level quality control, SNPs with MAFs  $<0.01$  and  $R^2$  values  $<0.3$  were excluded from the REML analysis. All analyses were adjusted for eigenvectors 1–20 from principal component analyses to account for possible confounding by population substructure within each cohort (22,43). Within the analysis, the component vectors were used as basic covariates to identify random genomic differences between genotyped data from cases and controls, in order to adjust statistical models for covariates accounting for possible population substructure. Summary statistics from these estimates were produced by every data set and were included in the meta-analyses.

The disease prevalence for PD was estimated from a general European ancestry populations identified by the literature. PD prevalence increases with age; therefore, the prevalence value was standardized for age and gender, and is specified here as 0.002 (4,8,17,18). To control for ascertainment, GCTA transformed the explained variance estimate from the observed scale ( $V(1)/V_p$ ) to the underlying scale ( $V(1)/V_{p\_L}$ ) to provide more robust heritability estimates.

Using PLINK version 1.07 (44), imputed SNPs were used to subset three basic data sets for each cohort. These included:

- (i) All imputed SNPs;
- (ii) All known SNPs  $\pm 1$  MB located within a region identified by replicated GWASs as associated with PD (described in Table 5); and
- (iii) SNPs not located within  $\pm 1$  MB of a region associated with PD by GWASs. Twenty-seven highly significant and well-replicated SNPs were used to define the PD regions for the second and third data sets, as described in Table 5.

Heritability estimates were compared between three data sets, when using:

- (i) All SNPs simultaneously;
- (ii) SNPs identified by GWASs as located within regions associated with PD; and
- (iii) SNPs not identified by GWASs (i.e. ‘missing’ or potentially hidden heritability).

We stratified our analysis to compare estimates between early and late PD onset. This allowed for a comparison of



**Table 5.** Loci associated with PD based on GWASs

Gene name(s)	Primary SNP	chr	Position (bp)	Citation(s)
<i>GBA</i>	N370S/i4000416	1	153 451 576–153 472 258	Lill <i>et al.</i> (48); Do <i>et al.</i> (15)
<i>SYT11/RAB25</i>	chr1:154105678	1	154 105 678	Lill <i>et al.</i> (48); Nalls and colleagues (32)
<i>RAB7L1/PARK16</i>	rs708723–rs947211	1	204 019 288	Lill <i>et al.</i> (48); Plagnol and colleagues (34)
<i>SLC41A1</i>	rs823156	1	204 031 263	Do <i>et al.</i> (15)
<i>ACMSD</i>	rs6710823	2	135 308 851	Nalls and colleagues (28)
<i>STK39</i>	rs2102808, rs2390669	2	168 800 188–168 825 271	Lill <i>et al.</i> (48); Nalls and colleagues (32)
<i>NMD3</i>	rs34016896	3	34 016 896	Plagnol and colleagues (34)
<i>MCCC1/LAMP3</i>	rs10513789, rs11711441	3	184 242 767–184 303 969	Lill <i>et al.</i> (48); Do <i>et al.</i> (15)
<i>GAK</i>	chr4: 811311, rs6599389	4	911 311	Nalls and colleagues (32)
<i>DGKQ</i>	rs11248060	4	929 113–954 359	Lill <i>et al.</i> (48); Do <i>et al.</i> (15)
<i>STBD1</i>	rs6812193	4	6 812 193	Plagnol and colleagues (34)
<i>BST1</i>	rs11724635	4	15 346 199	Lill <i>et al.</i> (48); Nalls and colleagues (32)
<i>SCARB2</i>	rs6812193	4	77 418 010	Do <i>et al.</i> (15)
<i>SNCA</i>	rs356220, rs6532194	4	90 999 925–90 860 363	Lill <i>et al.</i> (48); Nalls and colleagues (32); Do <i>et al.</i> (15)
<i>HLA-DRB5</i>	chr6:32588205	6	32 588 205	Nalls and colleagues (32)
<i>GNMB</i>	rs156429	7	156 429	Plagnol and colleagues (34)
<i>FGF20</i>	rs591323	8	591 323	Plagnol and colleagues (34)
<i>MMP16</i>	chr8:89442157	8	89 442 157	Plagnol and colleagues (34)
<i>ITGA8</i>	rs7077361	10	15 601 549	Lill <i>et al.</i> (48)
<i>LRRK2</i>	rs1491942, rs34637584	12	38 907 075–39 020 469	Lill <i>et al.</i> 2011 (48); Nalls and colleagues (32); Do <i>et al.</i> (15)
<i>CCDC62/HIP1R</i>	rs10847864, rs12817488	12	121 862 247–121 892 551	Lill <i>et al.</i> (48); Nalls and colleagues (32)
<i>STX1B</i>	rs4889603	16	4 889 603	Plagnol and colleagues (34)
<i>SREBF1/RAI1</i>	rs11868035	17	17 655 826	Do <i>et al.</i> (15)
<i>MAPT/STH</i>	rs12185268, rs2942168	17	41 149 582–42 131 818	Lill <i>et al.</i> (48); Nalls and colleagues (32); Do <i>et al.</i> (15)
<i>RIT2/SYT4</i>	rs4130047	18	38 932 233	Do <i>et al.</i> (15)
<i>USP25</i>	rs2823357	21	15 836 776	Do <i>et al.</i> (15)

Gene names and primary locations are provided for 27 highly significant and well-replicated SNPs. These SNPs were used to define the PD regions for our subsetted data sets.

‘heritability’ estimates for PD dependent on age at onset. REML analyses were repeated for GER, US and NL cohorts to compare early-onset ( $\leq 55$  years) cases and controls, and for GER, US, NL, UK, FR and ICE cohorts to compare late-onset cases ( $> 55$  years) and controls. In addition to cohort-level analyses, a meta-analysis was conducted to combine summary statistics.

We performed a random-effects meta-analysis using R 2.15. This produced powerful phenotypic variance and heterogeneity estimates across cohorts (Table 2). To quantify heterogeneity across cohorts, Cochran’s  $Q$  statistic was estimated using R. Cochran’s  $Q$  is based on a  $\chi^2$  distribution, and is calculated as the weighted sum of squared differences between each effect of the individual study and the pooled effect across studies (45–47):

$$Q = \sum w(\hat{E}E_C)^2, \quad E_C = \sum wE / \sum w.$$

This statistic does not inform the type or cause of heterogeneity, only of its presence or absence.

In order to quantify the impact of coverage of microarray genotyping on our heritability estimates, we determined the number of rare variants imputed into each cohort. Variants with an MAF of  $< 1\%$  were considered rare.

## WEB RESOURCES

R Development Core Team (2011). R: A Language and Environment for Statistical Computing. R Foundation for Statistical

Computing, Vienna, Austria, ISBN 3-900051-07-0, <http://www.R-project.org/>.

1000 Genomes Project: A Deep Catalog of Human Genetic Variation (2008–2010), <http://www.1000genomes.org/>.

MACH: MiniMac (2011), <http://genome.sph.umich.edu/wiki/Minimac>.

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## APPENDIX

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